A NOTE ON THE ESTIMATION OF EFFECTIVE NUMBER OF FACTORS BY MEANS OF PROBABILITY-GENERATING FUNCTIONS

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(Received December 2, 1965)

By a factor we shall mean any gene-pair or locus which is responsible for a particular character of a certain organism. Now consider a metric character of polygenic nature in a certain organism and consider a crossing between two pure line L_1 and L_2 of the organism. Suppose that there are n independent contributing factors in the cross, so that we can designate the genotypes of L_1 and L_2 respectively by $a_1a_1a_2a_2...a_na_n$ and $A_1A_1A_2A_2...A_nA_n$, L_1 being considered as having a smaller genotypic value in comparison with L_2 . Let D denote the difference between the genotypic values of L_1 and L_2 . Assuming that there is no interaction between the contributing factors, we then proceed to the estimation of n in accordance with the following two models:

Model I: Assume that there is no dominance and that each of the A genes has an additive and equal effect e; i.e.,

$$(1) e = \frac{D}{2n}.$$

Then, for some r such that $0 \le r \le 2n$, the relative frequency of an F_2 individual having r A genes and 2n - r a genes in its genotype will be

$${2n \choose r} \left(\frac{1}{2}\right)^{2n}.$$

The probability-generating function for the distribution of A genes is then given as follows:

(3)
$$P_{N}(t) = \left(\frac{1}{2}\right)^{2n} \sum_{r=0}^{2n} {2n \choose r} t^{r} = \left[\frac{1}{2}(t+1)\right]^{2n}$$

where N is the random variable representing the number of A genes, and t is a dummy variable. The average number of A genes in the F_2 population is

(4)
$$E(N) = \frac{d}{dt} P_{N}(t) \Big|_{t=1} = n.$$

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The first factorial moment of N is

(5)
$$E(N(N-1)) = \frac{d^{2}}{dt^{2}} P_{N}(t) \Big|_{t=1}$$

$$= \left(\frac{1}{2}\right)^{2n} \sum_{r=2}^{2n} r(r-1) {2n \choose r}$$

$$= \frac{n(2n-1)}{2},$$

whence it follows that

(6)
$$\operatorname{var}(N) = E(N(N-1)) + E(N) - E^{2}(N)$$

= $\frac{n}{2}$.

Hence the genetic variance, denoted by Vg, of F2 is given by

(7)
$$V_g = \text{var}(eN) = \text{var}\left(\frac{D}{2n} \cdot N\right)$$
$$= \frac{D^2}{8n}.$$

Thus n can be expressed in terms of D and V_g , as follows:

$$n = \frac{D^2}{8V_g}.$$

Model II: Assume that there is dominance and that each gene-pair AA or Aa has an additive and equal dominance effect

$$(9) d=2e=\frac{D}{n}.$$

Then, for some k satisfying $0 \le k \le n$, the relative frequency of an F_2 individual having k gene-pairs of the form AA or Aa and n-k gene-pairs of the form aa in its genotype is given by

$$\binom{n}{k} \left(\frac{3}{4}\right)^k \left(\frac{1}{4}\right)^{n-k},$$

and the probability-generating function for the distribution of the gene-pair AA or Aa will be

$$P_{M}(t) = \sum_{k=0}^{n} {n \choose k} \left(\frac{3}{4}\right)^{k} \left(\frac{1}{4}\right)^{n-k} t^{k} = \left[\frac{1}{4}(3t+1)\right]^{n}$$

where M is the random variable representing the number of gene-pairs of the form AA or Aa. The average number of such gene-pairs in the F_2 population is

(11)
$$E(M) = \frac{d}{dt} P_{M}(t) \Big|_{t=1}$$
$$= \frac{3n}{4}.$$

The first factorial moment of M is

(12)
$$E(M(M-1)) = \frac{d^{2}}{dt^{2}} P_{M}(t) \Big|_{t=1}$$
$$= \frac{9n(n-1)}{16}.$$

It then follows that

(13)
$$\operatorname{var}(M) = E(M(M-1)) + E(M) - E^{2}(M)$$

$$= \frac{3n}{16}.$$

Hence we obtain

(14)
$$V_g = \text{var} (dM) = \text{var} \left(\frac{D}{n} \cdot M \right)$$
$$= \frac{3D^2}{16n},$$

and so we can solve Equation (14) for n as follows:

(15)
$$n = \frac{3D^2}{16V_g}.$$

In practice, the parameter D may be estimated by the difference between the sample means of the L_1 and L_2 populations, denoted respectively by $\bar{L_1}$ and $\bar{L_2}$. Thus, denoting the estimate of D by \hat{D} , we have

$$\hat{\mathbf{D}} = \bar{\mathbf{L}}_2 - \bar{\mathbf{L}}_1.$$

Furthermore, we can, under the assumption of no interaction between the genotypes and the environmental effects and of no epistasis between the genepairs of the genotypes, decompose the variance of F_2 population, denoted by V_{F_2} as follows:

$$V_{F_2} = V_g + V_e,$$

where

$$V_g$$
 = Genetic variance,

and

V_e = Environmental variance.

Since all of the F_1 individuals have the same genotype, it is reasonable to express $V_{\mathfrak{s}}$ by the variance of the F_1 population, denoted by V_{F_1} . Thus we can estimate $V_{\mathfrak{s}}$ by

$$\hat{V}_{g} = \hat{V}_{F_{2}} - \hat{V}_{F_{1}},$$

where \hat{V}_g , \hat{V}_{F_2} , and \hat{V}_{F_1} are the estimates of V_g , V_{F_2} , and V_{F_1} respectively. Finally, the substitution of Equations (16) and (18) into Equations (8) and (15) leads respectively to

(19)
$$\hat{n} = \frac{(\bar{L}_2 - \bar{L}_1)^2}{8(\hat{V}_{F_2} - \hat{V}_{F_1})},$$

and

(20)
$$\hat{n} = \frac{3(\bar{L}_2 - \bar{L}_1)^2}{16(\hat{V}_{F_2} - \hat{V}_{F_1})},$$

where \hat{n} denotes the estimate of n.

介紹利用機率母函數以估算有效因子 總數的一個方法

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如果一個育種計劃的目標是在於改進某種作物、家畜、或家禽的定量形質 (metric character),則立計劃之前似乎應該考慮到究竟有多少有效的因子控制着該形質 ,如是才能預計育種的年限和按排育種的步驟。通常有效因子的總數越大,遺傳的效率越低,育種的年限也須要增長,同時在步驟方面也得根據集團育種的原理加以調整。本文於此介紹一個估算有效因子總數的方法,主要的構想是把 F_2 集團中每一因子型所含有的有效因子之數看作一個機率變數,藉以界定該變數的機率母函數並藉以探討各種可能的因子型在 F_2 集團中的分佈狀態,而利用 F_1 及 F_2 集團的觀察資料以估算控制該定量形質的有效因子之總數。

